

# UnderStandingAmericaStudy

WEIGHTING PROCEDURE, FEBRUARY 2023 - PRESENT



USC Dornsife Center for Economic and Social Research

**Contents**

- Introduction ..... 3
- 1. Sampling..... 3
  - 1.1. Respondents with a weight of zero ..... 5
- 2. Weighting..... 6
  - 2.1. Step 1: Base Weights..... 6
  - 2.2. Step 2: Poststratification Weights ..... 7
  - 2.3. Categorization and Imputation of Variables ..... 8
  - 2.4. Raking/Trimming Algorithm..... 10
  - 2.5. Final Poststratification Weights ..... 12
  - 2.6. Default Weights ..... 12
  - 2.7. Custom Weights ..... 15
  - 2.8. Weighting Output ..... 16

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## INTRODUCTION

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This document provides details of the weighting procedures and benchmark distributions used to create final sample weights for data sets collected by the Center for Economic and Social Research’s Understanding America Study internet panel.<sup>1</sup> The weighting procedure described in this document was implemented starting February 2023 and is currently in use.

Please consult the weighting procedures archive ([here](#)) for details about past weighting procedures adopted by the UAS.

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## 1. SAMPLING

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In this section, we provide a summary of UAS’s sampling procedures as background for our weighting protocol. For a full description of the UAS sampling and recruitment procedures, please check the UAS website at [uasdata.usc.edu](http://uasdata.usc.edu).

The UAS is a nationally representative panel of U.S. households recruited through Address Based Sampling (ABS). Eligible individuals are all adults in the contacted household aged 18 and older.

Sampling in the UAS is carried out in batches. There are currently 34 batches targeting either the U.S. population at large (batches 1, 5-12, 17, 20-28, 30-32, and 34) or specific subsets of it, such as the population of Native Americans (batches 2, 3, and 33), California residents (batches 15 and 16), and Los Angeles County residents (batches 13-14, 18-19, and 29).<sup>2</sup> Table 1 below lists all the UAS recruitment batches as of April 2024 and their corresponding sample frames and sampling methods.

### Sampling method: SRS

Batch 1 is a simple random sample (SRS) from the list of all individuals in the ASDE Survey Sampler database. Batches 22-23 are simple random samples of all U.S. addresses provided by the vendor

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<sup>1</sup> Mick Couper and Jon Krosnick have provided insightful and valuable comments throughout the development of the UAS weighting procedure.

<sup>2</sup> Although 34 recruitment batches are listed in UAS datasets, the actual number of recruitment batches is 32. Batches 24 and 26 were split in two (24&25 and 26&27, respectively) to carry out recruitment experiments.

Marketing Systems Group (MSG). Batch 4 is a simple random sample of all addresses listed on birth certificates issued in Los Angeles County in the years 2009-2012 in a limited set of zip codes. Batches 29 is a simple random sample of all Los Angeles County addresses provided by the vendor Marketing Systems Group (MSG). Batch 33 is a simple random sample of addresses in areas with high a concentration of Native Americans.

#### Sample method: Adaptive1

Batches 2-21 are based on a two-stage sample design, in which zip codes are drawn first (within a pre-defined sample frame), and then households are randomly drawn from the sampled zip codes. More precisely, these batches are selected using an adaptive sampling algorithm, which we refer to as Adaptive1. This algorithm allows to refresh the panel in such a way that its demographic composition moves closer to the population composition. Specifically, before sampling an additional batch, the algorithm computes the unweighted distributions of specific demographic characteristics (e.g., sex, age, marital status and education) in the UAS at that point in time. It then assigns to each zip code a non-zero probability of being drawn, which is an increasing function of the degree of “desirability” of the zip code. The degree of desirability is a measure of how much, given its population characteristics, a zip code is expected to move the current distributions of demographics in the UAS towards those of the U.S. population. For example, if at a particular point in time the UAS panel underrepresents females with high school degree, zip codes with a relatively high proportion of females with high school degree receive a higher probability of being sampled. This sampling algorithm is implemented iteratively.

That is, after selecting a zip code, the distributions of demographics in the UAS are updated according to the expected contribution of this zip code towards the panel’s representativeness, updated measures of desirability are computed and new sampling probabilities for all other zip codes are defined. Such procedure provides a list of zip codes to be sampled. The implementation of this algorithm implies that the marginal probability of drawing each zip code depends on the composition of the UAS panel at a particular point in time, but also on the unknown response probabilities of selected households in that zip code. Hence, the marginal probability of drawing each zip code is not known ex ante and cannot be used to construct design weights. The weighting procedure corrects for the unequal sampling probabilities generated by this adaptive sampling algorithm.

#### Sample method: Adaptive2

Batches 24-28, 30-32, and 34 are random samples of all U.S. addresses provided by MSG, with selection probabilities generated by a second adaptive sampling algorithm, which we refer to as Adaptive2. This second sampling approach involves the following two steps. First, we combine auxiliary information attached to the post office delivery sequence files provided by MSG and information about population characteristics at the Census tract level to compute the probability of belonging to a certain population stratum for each potential invitee. Second, we assign sample inclusion probabilities so that, in expectation, the final sample of selected invitees has a desired distribution of demographic characteristics. Such desired distribution is defined in order to increase the representation of under-represented segments of the population in the UAS (e.g., individuals with high school or less) and/or to obtain sufficient sample sizes to detect meaningful differences between groups (e.g., differences in socio-economic status and health outcomes between racial/ethnic groups). The inverse of the inclusion probabilities generated by this second adaptive sampling algorithm represents the individual-level base weight that accounts for differential selection probabilities across sampled individuals.

**Table 1: UAS Recruitment Batches**

Batch	Sample Frame	PSU	Method
1	U.S.	Address	SRS
2-3	Areas with high concentrations of Native Americans	Zip Code	Adaptive1
4	Los Angeles County (birth certificate list sample)	Address	SRS
5-12	U.S.	Zip Code	Adaptive1
13-14	Los Angeles County	Zip Code	Adaptive1
15-16	California	Zip Code	Adaptive1
17	U.S.	Zip Code	Adaptive1
18-19	Los Angeles County	Zip Code	Adaptive1
20-21	U.S.	Zip Code	Adaptive1
22-23	U.S.	Address	SRS
24-25	U.S.	Address	Adaptive2
26-28	U.S.	Address	Adaptive2
29	Los Angeles County	Address	SRS
30-32	U.S.	Address	Adaptive2
33	Areas with high concentrations of Native Americans	Address	SRS
34	U.S.	Address	Adaptive2

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### 1.1. Respondents with a weight of zero

Recruitment batches 2-3 and 33 target the Native American population. More precisely, the sample frame for these batches is areas with a high concentration of Native Americans. Even though non-Native Americans contacted within batches 2 and 3 were not eligible to become panel members, some were accidentally invited to join the UAS. Because we are unable to assign a probability to this happening, these panel members receive a weight of zero. Batch 33 only included 756 addresses. It was recruited for the specific purpose of evaluating participation rates within Native American reserves and assessing the cost of recruiting a sizeable sample from these areas. Because of the “exploratory” nature of batch 33, all respondents within it receive a weight of zero.

As mentioned above, recruitment batch 4 was a simple random sample of addresses listed on birth certificates issued in Los Angeles County in the years 2009-2012 in a limited set of zip codes. Because of the highly specific nature of this subsample, we do not provide weights for members recruited within this batch and assign to all of them a weight of zero.

Thus, **we provide weights for respondents in all batches listed in Table 1, except for non-Native American respondents in batches 2 and 3 and all respondents in batches 33 and 4.**

Non-Native American respondents in batches 2 and 3 and respondents in batches 33 and 4 are often included in UAS survey data sets. In this case, **they all have a sample weight equal to zero.** Hence, they are dropped whenever weighted statistics are computed. Data users should decide whether to retain these respondents for analyses that do not require the use of sample weights.

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## 2. WEIGHTING

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In the UAS, sample weights are survey-specific. They are provided with each UAS survey and, unless otherwise indicated, are meant to make each survey data set representative of the U.S. population with respect to a pre-defined set of socio-demographic variables. Sample weights are constructed in two steps. In the first step, a *base weight* is created to account for unequal probabilities of sampling UAS members generated by the adaptive sampling algorithms described above. In the second step, *final post-stratification weights* are generated to correct for differential non-response rates and to bring the final survey sample in line with the reference population as far as the distribution of key variables of interest is concerned.

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## 2.1. Step 1: Base Weights

### *Base weights for members sampled via Adaptive1*

When computing base weights for batches using Adaptive1, the unit of analysis is a zip code. We estimate a logit model for the probability that a zip code is sampled as a function of its characteristics, namely Census region, urbanicity, population size, as well as sex, race, age, marital status, and education composition. Estimation is carried out on an American Community Survey (ACS) file that contains 5-year average characteristics at the zip code level.<sup>3</sup> The outcome of this logit model is an estimate of the marginal probability of a zip code being sampled, which, given the implementation of the adaptive sampling algorithm Adaptive1 described above, is not known *ex-ante*.

We indicate by  $\pi_k$  the logit estimated probability of sampling zip code  $k$ . The probability of sampling household  $h$  after drawing zip code  $k$  is the ratio of the number of households sampled divided by the number of households in the zip code. We indicate this by  $\pi_{h|k}$ . Hence, the marginal probability that household  $h$  from zip code  $k$  is sampled is  $\pi_{hk} = \pi_{h|k} \times \pi_k$ .

The base weight is a zip code level weight defined as:

$$w_{hk}^{base} = \Lambda \times \frac{1}{\pi_{hk}}$$

where the constant  $\Lambda$  is chosen such that the sum of the base weights is equal to the number of sampled households. A comprehensive discussion of how base weights for UAS members sampled via Adaptive 1 are computed is provided in Angrisani et al. (2020), available [here](#).

UAS members sampled via Adaptive 1 are assigned a base weight, computed as described above, depending on the zip code where they reside at the time of recruitment.

### *Base weights for members sampled via Adaptive2*

The adaptive sampling algorithm Adaptive2 assigns specific inclusion probabilities to all invitees depending on their demographic characteristics and on the desired distribution of demographic

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<sup>3</sup> Strictly speaking, all files from the U.S. Census Bureau use "zip code tabulation area" (zcta), which is based on, but not identical to, USPS's definition of zip codes. We ignore the distinction between the two.

characteristics in the entire UAS that should be achieved with each refresher batch. The inverse of these inclusion probabilities constitutes the individual-level base weight that accounts for differential selection probabilities across UAS members sampled via Adaptive2.

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## 2.2. Step 2: Poststratification Weights

The execution of the sampling process for a survey is typically less than perfect. Even if the sample of panel members invited to take a survey is representative of the population along a series of dimensions, the sample of actual respondents may exhibit discrepancies because of differences in response rates across groups and/or other issues related to the fielding time and content of the survey. A second layer of weighting is therefore needed to align the final survey sample to the reference population as far as the distribution of key variables is concerned.

In this second step, we perform **raking weighting** (also known as iterative marginal weighting), starting from the base weights,  $w_{hk}^{base}$ , described in the previous section. With this, we assign poststratification weights to survey respondents such that the weighted distributions of specific socio-demographic variables in the survey sample match their population counterparts (benchmark or target distributions).

The benchmark distributions against which UAS surveys are weighted are derived from the Basic Monthly Current Population Survey (CPS).<sup>4</sup> We use the 6 most recent available monthly CPS at the time a UAS survey is completed. This ensures a minimum gap between the period of survey completion and the period benchmark distributions refer to.

Unless otherwise required by the survey's aims and specified in the sample selection process (e.g., a survey targeting only Los Angeles County residents or individuals over the age of 65), the reference population for UAS surveys is the U.S. population of adults age 18 or older, excluding institutionalized individuals and military personnel.

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## 2.3. Categorization and Imputation of Variables

For post-stratification weighting purposes, we use demographic information taken from the most recent *My Household* survey, which is answered by all active UAS members every quarter. All

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<sup>4</sup> We rely on IPUMS-CPS, University of Minnesota, <https://ipums.org/>.



socio-demographic variables in the *My Household* survey are categorical, but some, such as age, education, and income, take values in a relatively large set. We recode all socio-demographic variables considered for poststratification into new categorical variables with no more than 5 categories. The aim of limiting the number of categories is to prevent these variables from forming strata containing a very small fraction of the sample (less than 5%), which may cause sample weights to exhibit considerable variability.

The list of all recoded categorical variables considered for poststratification is reported in Table 2.

Before implementing the poststratification weighting procedure, we employ the following imputation scheme to replace missing values of recoded socio-demographic variables.

- Gender is obtained from administrative records.
- When age is missing, the age range available in the My Household survey is used to impute age categories. If the age range is also missing, age categories are imputed using gender-specific sample mode.
- Once age categories have been imputed (if missing), the variable with the fewest missing values is the first one to be imputed by means of a regression featuring gender and the age categories as regressors. This newly imputed variable is then added to the set of regressors to impute the variable with the second smallest number of missing values. The procedure continues in this fashion until the variable with the most missing values is imputed using information on all other available socio-demographic variables.
- For binary indicators, such as born in the US and US citizen, missing values are imputed using a logistic regression. For ordered categorical variables, such as education, household composition, and household income, missing values are imputed using an ordered logistic regression. For unordered categorical variables, such as marital status, race-ethnicity, and work status, missing values are imputed using a multinomial logistic regression. Census regions are never missing, as they are obtained from respondents' zip codes of residence.

Each UAS survey data set with sample weights contains a binary variable (*imputation\_flag*) indicating whether any of the recoded socio-economic variables used for poststratification has been imputed or taken from UAS administrative records not available to data users. The mean of the imputation flag across UAS surveys is typically between 0.01 and 0.015.

Table 2: List of Recoded Categorical Variables for Poststratification

Recoded Variable	Categories
<i>Gender</i>	1. Male; 2. Female
<i>Age</i>	1. 18-39; 2. 40-49; 3. 50-59; 4. 60+
<i>Born in the US</i>	0. No; 1. Yes
<i>US citizen</i>	0. No; 1. Yes
<i>Education</i>	1. High School or Less; 2. Some College; 3. Bachelor or More
<i>Native American</i>	0. No; 1. Yes
<i>Race-ethnicity</i>	1. White; 2. Black; 3. Others; 4. Hispanic; 5. Native American
<i>Census region (aug)*</i>	1. Northeast; 2. Midwest; 3. South; 4. West, excl CA; 5. CA, excl LAC; 6. LAC
<i>Marital status</i>	1. Married; 2. Separated/Divorced/Widowed; 3. Never Married
<i>Work status</i>	1. Working; 2. Unemployed; 3. Retired; 4. On leave, Disabled, Other
<i>Household composition</i>	1. 1 Member; 2. 2 Members; 3. 3 or 4 Members; 4. 5 or More Members
<i>Household income</i>	1. <\$30,000; 2. \$30,000-\$59,999; 3. \$60,000-\$99,999; 4. \$100,000+

\* These are obtained from the respondent's zip code of residence.

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## 2.4. Raking/Trimming Algorithm

We adopt a **raking algorithm** to generate poststratification weights. This procedure involves the comparison of target population relative frequencies and actually achieved sample relative frequencies on a number of socio-demographic variables independently and sequentially. More precisely, starting from the base weights, at each iteration of the algorithm weights are proportionally adjusted so that the distance between survey and population marginal distributions of each selected socio-demographic variable (or raking factor) decreases. The algorithm stops when survey and population distributions are perfectly aligned. A maximum of 50 iterations is allowed for perfect alignment of survey and population distributions to be achieved. If the process does not converge within 50 iterations, no sample weights are returned and attempts using different raking factors are made.

Our raking algorithm trims extreme weights in order to limit variability and improve efficiency of estimators. We follow the general weight trimming and redistribution procedure described by Valliant, Dever and Kreuter (2013).<sup>5</sup>

Specifically, we define  $N = N_w + N_{nw}$  the total sample size, where  $N_w$  is the number of respondents who receive a weight, and  $N_{nw}$  is the number of respondents with a pre-assigned weight of zero (non-Native Americans in batches 2 and 3; all respondents in batch 4). Indicating with  $w_i^{rak}$  the raking weight for respondent  $i = 1, \dots, N_w$ , and with  $\bar{w}^{rak} = \frac{1}{N_w} \sum_{i=1}^{N_w} w_i^{rak}$  the sample average of raked weights,

- I. We set the lower ( $L$ ) and upper ( $U$ ) bounds on weights equal to the 10th and 90th percentile of the  $w_i^{rak}$  distribution, respectively. While there is no consensus on which threshold should be used for trimming, these are among those often mentioned in the literature and adopted by other surveys (Battaglia et al., 2009).<sup>6</sup>
- II. We reset any weights smaller than the lower bound to  $L$  and any weights greater than the upper bound to  $U$ :

$$w_i^{trim} = \begin{cases} L & w_i^{rak} \leq L \\ w_i^{rak} & L < w_i^{rak} < U \\ U & w_i^{rak} \geq U \end{cases}$$

- III. We compute the amount of weight lost by trimming as  $w^{lost} = \sum_{i=1}^{N_c} (w_i^{rak} - w_i^{trim})$  and distribute it equally among the respondents whose weights are not trimmed.
- IV. If these new weights are all within the interval  $[L,U]$ , no further adjustment is performed. If any of these new weights are outside the interval  $[L,U]$ , the trimming procedure is

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<sup>5</sup> Valliant, R., Dever, J. A., and Kreuter F., (2013) *Practical Tools for Designing and Weighting Survey Samples*. Springer, New York.

<sup>6</sup> Battaglia, M. P, Izrael, D., Hoaglin, D. C., and Frankel M. R., (2009) "Practical Considerations in Raking Survey Data." *Survey Practice*, 2009 (June). <http://surveypractice.org/2009/06/29/raking-survey-data/>.

repeated iteratively until all weights are within the interval [L,U] or until the maximum number of 50 iterations is reached.

While raking weights can match population distributions of selected variables, trimmed weights typically do not. We, therefore, iterate the raking algorithm and the trimming procedure until post-stratification weights are obtained that respect the weight bounds and align sample and population distributions of selected variables. This procedure stops after 50 iterations if an exact alignment respecting the weight bounds cannot be achieved. In this case, the raked weights will ensure an exact match of (weighted) survey relative frequencies to their population counterparts, but the weights will not be within the pre-determined bounds.

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## 2.5. Final Poststratification Weights

Indicate by  $w_i^{post}$  the final poststratification weight for respondent  $i$ , obtained by applying the raking algorithm to the base weights and after iterating the raking/trimming procedure as described above. Each weighted UAS survey data set includes final poststratification weights relative to their sample mean. Formally, this is

$$w_i^{final} = \frac{w_i^{post}}{\left(\frac{1}{N_w} \sum_{j=1}^{N_w} w_j^{post}\right)}$$

For the  $N_w$  respondents who receive a weight, and  $w_i^{final} = 0$  for the  $N_{nw}$  respondents who do not. Hence, relative final poststratification weights sum to the size of the sample of respondents who receive a weight ( $N_w$ ) and average to 1 within that sample.

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## 2.6. Default Weights

Poststratification weights for UAS surveys including all batches are generated using the following set of raking factors (as defined in Table 2):

- ❖ *Gender*
- ❖ *Race-ethnicity*
- ❖ *Age*

❖ *Education*

❖ *Census region (aug)*

For UAS surveys including only batches targeting the U.S. population (batches 1, 5-12, 17, 20-28, 30-32, and 34), the set of raking factors includes gender, race-ethnicity (excluding Native Americans), age, education, and Census region (Northeast, Midwest, South, West).

We have carried out extensive testing and concluded that raking weights produced by this combination of factors perform well across different dimensions. In particular, they exhibit moderate variability, thereby leading to better precision of weighted estimates, and allow matching the distributions of variables not used as raking factors in a satisfactory manner, thereby improving overall representativeness. Our Monte Carlo studies have shown that these desirable properties are robust across the typical sample sizes of UAS surveys.

For UAS surveys currently in the field, default weights can be obtained by sending a request to [uas-weights-l@mymailists.usc.edu](mailto:uas-weights-l@mymailists.usc.edu).

For completed surveys, the data set with default weights is available for download on the [UAS webpage](#).

Table 3 shows the distributions of the default raking factors in the entire UAS sample as of April 2024. The demographic information used in this exercise is obtained from the My Household survey as of April 1<sup>st</sup>, 2024 (each UAS member is asked to update their My Household survey every quarter).

In column 1 of Table 3, we report unweighted percentages; in column 2, the percentages after applying final weights. By construction, these match population benchmarks, which are taken from the Basic Monthly Current Population Survey (months from July 2023 to December 2023). The unweighted sample size (excluding respondents from a weight of zero) is 14,269; the effective sample size using the final poststratification weights is 7,848 (55% of 14,269).

**Table 3: Unweighted and Weighted Distributions of Raking Factors (%)**

Variable	Unweighted	Final weights*
Male	39.25	48.81
<i>Race-ethnicity</i>		
White	58.91	61.10
Black	11.75	12.06
Other	9.29	7.46
Hispanic	16.20	16.91
Native American	3.85	2.47
<i>Age (years)</i>		
18–39	30.24	37.47
40–49	19.33	15.92
50–59	17.75	15.73
60+	32.68	30.89
<i>Education</i>		
High school or less	20.34	37.75
Some college	33.56	26.45
Bachelor or more	46.10	35.80
<i>Census region (aug)</i>		
Northeast	14.33	17.39
Midwest	19.23	20.46
South	31.99	38.55
West, excl CA	8.49	11.97
CA, excl LAC	11.68	8.68
LAC	14.27	2.95

\*Weighted percentages using final weights match their population counterparts by construction, given the convergence of the raking/trimming algorithm.

Table 4 shows the distributions of some of the demographics included in the My Household survey (transformed as described in Table 2) that are not used as raking factors. In column 1, we report unweighted percentages; in column 2, the percentages after applying the final weights; and in column 3, the percentages in the population (taken from the CPS).

**Table 4: Unweighted and Weighted Distributions of Selected Variables Not Used as Raking Factors (%)**

Variable	Unweighted	Final weights	Population
<i>Household size</i>			
1 member	19.78	17.30	19.00
2 members	41.08	40.93	34.26
3 or 4 members	29.34	30.72	33.66
5 or more members	9.80	11.06	13.08
<i>Household income</i>			
<\$30,000	22.37	25.55	15.44
\$30,000–\$59,999	22.15	23.51	22.95
\$60,000–\$99,999	22.11	21.87	23.39
≥\$100,000	33.37	29.08	38.23
<i>Work status</i>			
Working	60.08	58.93	59.55
Retired	21.63	20.09	20.28
Other	18.29	20.98	20.17
<i>Employment type</i>			
Government	18.91	17.40	13.19
Private	69.10	71.67	76.58
Self-employed	11.98	10.93	10.23

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## 2.7. Custom Weights

Data users can customize the weighting procedure and obtain weights that better suit the goals of their research and data analysis. Custom weights can be obtained by choosing which socio-demographic variables should be used by the raking algorithm to generate post-stratification weights. Raking can be performed on one-way marginals, by matching population distributions of single socio-demographic variables, such as gender or education, as well as on two-way marginals, by matching the distributions of interaction variables, such as gender × education. The preferred set of raking factors may feature both single and interaction variables, such as, for instance, race-ethnicity and gender × education. The use of two-way marginals corrects for discrepancies between distributions referring to specific subgroups that would not be accounted for by using one-way marginals alone.

Data users can also request custom weights by targeting different reference populations both in terms of characteristics (e.g., the population of individuals older than 65) and geographical location (e.g., the population of California or Los Angeles County residents).

Custom weights requests should be sent to [uas-weights-l@mymailists.usc.edu](mailto:uas-weights-l@mymailists.usc.edu), alongside with the preferred set of raking factors. This set is **limited to the variables listed in Table 2 and to a maximum of 6 variables**. (single variables, interaction variables or a combination of both). Restrictions on the number and type of raking factors are imposed to ensure convergence of the algorithm and to reduce weight variability.

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## 2.8. Weighting Output

Each weighted UAS survey data set includes the following variables:

- *imputation\_flag*  
A binary variable indicating whether any of the variables used for poststratification has been imputed.
- *cps*  
A variable indicating the CPS monthly surveys used to obtain benchmark distributions for poststratification.
- *final\_weight*  
Relative final poststratification weights ensuring representativeness of the survey sample with respect to key pre-selected demographic variables.